(5 points)

(5 points)

1.  $Z_t$  is a sequence of independent standard normal random variables.

Let 
$$X_t = \begin{cases} Z_t & t \text{ is odd} \\ aZ_{t-1}^2 + b & t \text{ is even} \end{cases}$$

(1) Find the mean function of 
$$X_t$$
.

$$E(X_t) = \begin{cases} E(Z_t) = 0 & t \text{ is odd} \\ E(aZ_{t-1}^2 + b) = a + b & t \text{ is even} \end{cases}$$

(2) Find the autocovariance function of  $X_t$ .

$$\gamma(0) = \operatorname{var}(X_t) = \begin{cases} \operatorname{var}(Z_t) &= 1 \quad t \text{ is odd} \\ \operatorname{var}(aZ_{t-1}^2 + b) &= 2a^2 \quad t \text{ is even} \end{cases}$$

$$\gamma(1) = \begin{cases} \operatorname{cov}(X_{2t}, X_{2t+1}) &= \operatorname{cov}(aZ_{2t-1}^2 + b, Z_{2t+1}) &= 0 \\ \operatorname{cov}(X_{2t-1}, X_{2t}) &= \operatorname{cov}(Z_{2t-1}, aZ_{2t-1}^2 + b) &= 0 \end{cases}$$

$$\gamma(-1) = \begin{cases} \operatorname{cov}(X_{2t}, X_{2t-1}) &= \operatorname{cov}(aZ_{2t-1}^2 + b, Z_{2t-1}) &= 0 \\ \operatorname{cov}(X_{2t+1}, X_{2t}) &= \operatorname{cov}(Z_{2t+1}, aZ_{2t-1}^2 + b) &= 0 \end{cases}$$

$$\gamma(h) = 0 \text{ when } |h| > 1.$$

(3) Is 
$$X_t$$
 a strictly stationary time series? (5 points)

No. If  $X_t$  is strictly stationary, then  $X_1 = Z_1 \sim N(0, 1^2)$  and  $Z_2 = aZ_1^2 + b \sim a\chi^2(1) + b$  would have had the same distribution. But they do not.

(4) Is 
$$X_{2t+1}$$
 a strictly stationary time series? (5 points)

Yes.  $X_{2t+1} = Z_{2t+1}$  are iid  $N(0, 1^2)$  and hence is strictly stationary.

(5) Is 
$$X_{2t}$$
 a stationary time series? (5 points)

$$X_{2t} = aZ_{2t-1}^2 + b$$
 are iid  $a\chi^2(1) + b$  and hence is strictly stationary. Also  $var(X_{2t}) = 2a^2 < \infty$ . So  $E(X_{2t}^2) < \infty$ . Thus  $X_{2t}$  is stationary.

(6) Find conditions on 
$$a$$
 and  $b$  such that  $X_t$  is a white noise. (5 points)

If  $a = \frac{1}{\sqrt{2}}$  and  $b = -\frac{1}{\sqrt{2}}$ , then  $X_t$  represents an uncorrelated series with mean 0 and variance 1, and hence is a white noise.

(7) Under the conditions in (6) are 
$$X_t$$
 and  $X_{t+1}$  independent? (5 points)

No. 
$$X_2 = \frac{1}{\sqrt{2}} Z_1^2 - \frac{1}{\sqrt{2}} \sim \frac{1}{\sqrt{2}} \chi^2(1) - \frac{1}{\sqrt{2}}$$
.  
But  $X_2 | (X_1 = 1) = 0 \nsim \frac{1}{\sqrt{2}} \chi^2(1) - \frac{1}{\sqrt{2}}$ . So  $X_1$  and  $X_2$  are not independent.

(8) Under the conditions in (6) are 
$$X_t$$
 and  $X_{t+1}$  identically distributed? (5 points)

No. 
$$X_1 = Z_1 \sim N(0, 1)$$
, but  $X_2 = \frac{1}{\sqrt{2}}Z_1^2 - \frac{1}{\sqrt{2}} \sim \frac{1}{\sqrt{2}}\chi^2(1) - \frac{1}{\sqrt{2}}$ .

- 2.  $X_t$  is stationary with mean  $\mu_X$  and ACVF  $\gamma_X(h)$ .  $Y_t$  is stationary with mean  $\mu_Y$  and ACVF  $\gamma_Y(h)$ .  $X_t$  and  $Y_t$  are independent, i.e.,  $X_r$  and  $Y_t$  are independent.
  - (1) Is  $X_t + Y_t$  stationary? If yes, find its mean and ACVF. (10 points)

Yes  $X_t + Y_t$  is stationary with  $E(X_t + Y_t) = \mu_X + \mu_Y$  and  $\gamma_{X+Y}(h) = \text{cov}(X_t + Y_t, X_{t+h} + Y_{t+h}) = \gamma_X(h) + \gamma_Y(h)$ .

(2) Is  $X_t - Y_t$  stationary? If yes, find its mean and ACVF. (10 points)

Yet  $X_t - Y_t$  is stationary with  $E(X_t - Y_t) = \mu_X - \mu_Y$  and  $\gamma_{X-Y}(h) = \text{cov}(X_t - Y_t, X_{t+h} - Y_{t+h}) = \gamma_X(h) + \gamma_Y(h)$ .

(3) Is  $X_t \cdot Y_t$  stationary? If yes, find its mean and ACVF. (10 points)

Yes  $X_t \cdot Y_t$  is stationary with  $E(X_t \cdot Y_t) = \mu_X \cdot \mu_Y$  and  $\gamma_{XY}(h) = \cot(X_t Y_t, X_{t+h} Y_{t+h}) = E(X_t Y_t X_{t+h} Y_{t+h}) - E(X_t Y_t) E(X_{t+h} Y_{t+h})$   $= E(X_t X_{t+h}) E(Y_t Y_{t+h}) - \mu_X \mu_Y \mu_X \mu_Y$   $= [\gamma_X(h) + \mu_X^2] [\gamma_Y(h) + \mu_Y^2] - \mu_X^2 \mu_Y^2$   $= \mu_X^2 \gamma_Y(h) + \mu_Y^2 \gamma_X(h) + \gamma_X(h) \gamma_Y(h).$ 

- 3.  $X_1$ ,  $X_2$ ,  $X_4$ ,  $X_5$  are observed from  $X_t = \phi X_{t-1} + Z_t \sim AR(1)$  where  $Z_t \sim WN0$ ,  $\sigma^2$ ) and  $0 < |\phi| < 1$ .
  - (1) Find the best linear predictor  $P_2X_8$  of  $X_8$  using  $X_1$  and  $X_2$ . What is the corresponding mean squared error? (15 points)

 $X_t = \phi X_{t-1} + Z_t \sim \operatorname{AR}(1) \Longrightarrow E(X_t) = 0 \text{ and } \gamma(h) = \frac{\phi^h \sigma^2}{1 - \phi^2}.$   $\Gamma_2 = \operatorname{Cov}\left(\begin{pmatrix} X_1 \\ X_2 \end{pmatrix}\right) = \frac{\sigma^2}{1 - \phi^2} \begin{pmatrix} 1 & \phi \\ \phi & 1 \end{pmatrix} \text{ and } \gamma_2 = \operatorname{Cov}\left(\begin{pmatrix} X_1 \\ X_2 \end{pmatrix}, X_8 \right) = \frac{\sigma^2}{1 - \phi^2} \begin{pmatrix} \phi^7 \\ \phi^6 \end{pmatrix}$   $\mathbf{a}_2 = \Gamma_2^{-1} \gamma_2 = \begin{pmatrix} 1 & \phi \\ \phi & 1 \end{pmatrix}^{-1} \begin{pmatrix} \phi^7 \\ \phi^6 \end{pmatrix} = \begin{pmatrix} 0 \\ \phi^6 \end{pmatrix}. \text{ So } P_2 X_8 = \mathbf{a}_2' \begin{pmatrix} X_1 \\ X_2 \end{pmatrix} = \phi^6 X_2.$ The mean squared error:  $\gamma(0) - \mathbf{a}_2' \gamma_2 = \frac{\sigma^2}{1 - \phi^2} - (0, \phi^6) \frac{\sigma^2}{1 - \phi^2} \begin{pmatrix} \phi^7 \\ \phi^6 \end{pmatrix} = \frac{1 - \phi^{12}}{1 - \phi^2} \sigma^2.$ 

(2) Find the best linear estimate of the missing value  $X_3$  using  $X_4$  and  $X_5$ . Compute the mean squared error of this estimate. (15 points)

$$X_{t} = \phi X_{t-1} + Z_{t} \sim \operatorname{AR}(1) \Longrightarrow E(X_{t}) = 0 \text{ and } \gamma(h) = \frac{\phi^{h} \sigma^{2}}{1 - \phi^{2}}.$$

$$\Gamma_{2} = \operatorname{Cov}\left(\begin{pmatrix} X_{4} \\ X_{5} \end{pmatrix}\right) = \frac{\sigma^{2}}{1 - \phi^{2}} \begin{pmatrix} 1 & \phi \\ \phi & 1 \end{pmatrix} \text{ and } \gamma_{2} = \operatorname{Cov}\left(\begin{pmatrix} X_{4} \\ X_{5} \end{pmatrix}, X_{3} \right) = \frac{\sigma^{2}}{1 - \phi^{2}} \begin{pmatrix} \phi \\ \phi^{2} \end{pmatrix}$$

$$\mathbf{a}_{2} = \Gamma_{2}^{-1} \gamma_{2} = \begin{pmatrix} 1 & \phi \\ \phi & 1 \end{pmatrix}^{-1} \begin{pmatrix} \phi \\ \phi^{2} \end{pmatrix} = \begin{pmatrix} \phi \\ 0 \end{pmatrix}. \text{ So } P_{2} X_{3} = \mathbf{a}_{2}' \begin{pmatrix} X_{4} \\ X_{5} \end{pmatrix} = \phi X_{4}.$$
The mean squared error:  $\gamma(0) - \mathbf{a}_{2}' \gamma_{2} = \frac{\sigma^{2}}{1 - \phi^{2}} - (\phi, 0) \frac{\sigma^{2}}{1 - \phi^{2}} \begin{pmatrix} \phi \\ \phi^{2} \end{pmatrix} = \sigma^{2}.$