L13 Simultaneous Confidence Regions

- 1. Confidence regions in linear model Consider the model $Y = X\beta + \epsilon$, $\epsilon \sim N(0, \sigma^2 \Sigma)$ where $X \in \mathbb{R}^{n \times p}$ has full column rank.
 - (1) Confidence interval for σ^2

$$\left(\frac{\text{SSE}}{\chi^2_{\alpha/2}(n-p)}, \frac{\text{SSE}}{\chi^2_{1-\alpha/2}(n-p)}\right) \text{ is a } 1-\alpha \text{ confidence interval for } \sigma^2.$$

(2) Confidence region for $\theta = H\beta \in R^q$. The collection of all $\theta \in R^q$ satisfying

$$\frac{(\theta - H\widehat{\beta})'[H(X'\Sigma^{-1}X)^{-1}H']^{-1}(\theta - H\widehat{\beta})}{q \text{ MSE}} \le F_{\alpha}(q, n-p)$$

is a $1 - \alpha$ C.R. for $\theta = H\beta$ where $H \in \mathbb{R}^{q \times p}$ has full row rank. For example

$$\frac{(\beta - \widehat{\beta})'(X'\Sigma^{-1}X)(\beta - \widehat{\beta})}{p \text{ MSE}} \le F_{\alpha}(p, n - p)$$

is $1 - \alpha$ C.R. for $\beta \in \mathbb{R}^p$.

(3) Confidence interval for $\theta = l'\beta \in R$ $\theta \in l'\widehat{\beta} \pm \sqrt{F_{\alpha}(1, n-p)} \sqrt{\text{MSE } l'(X'\Sigma^{-1}X)l}$ is a $1 - \alpha$ C.I. for $l'\beta \in R$. In HW the above formula has been re-written as $\theta \in l'\widehat{\beta} \pm t_{\alpha/2}(n-p)S_{l'\widehat{\beta}}$.

2. Bonferroni simultaneous confidence regions

(1) An inequality

Let $A_i = [\theta_i \in B_i]$ be a random event, i = 1, ..., k, such that $P(A_i) \ge 1 - \frac{\alpha}{k}$. Then $P(A_1 \cap \dots \cap A_i) \ge 1 - \alpha$. **Proof.** $P(A_1 \cap \dots \cap A_k) = 1 - P(A_1^c \cup \dots \cup A_k^c) \ge 1 - [P(A_1^c) + \dots + P(A_k^c)]$ $= 1 - [1 - P(A_1) + \dots + 1 - P(A_k)]$ $= 1 - k + [P(A_1) + \dots + P(A_k)] \ge 1 - k + (1 - \frac{\alpha}{k}) k$ $= 1 - k + k - \alpha = 1 - \alpha$.

(2) Bonferroni method

If $\theta_i \in B_i$ is a CR for θ_i with confidence coefficient $1 - \frac{\alpha}{k}$, i = 1, ..., k, then $\theta_i \in B_i$, i = 1, ..., k, are simultaneous CRs for θ_i , i = 1, ..., k, with overall confidence coefficient $1 - \alpha$.

Ex1: Find formulas for confidence intervals for $\beta_1, ..., \beta_p$ with overall confidence coefficient 80%.

 $\beta_i \in \widehat{\beta}_i \pm t_{0.20/(2p)}(n-p) S_{\widehat{\beta}_i}, i = 1, ..., p$ are simultaneous CIs for $\beta_i, i = 1, ..., p$, with overall confidence coefficient 80%.

Comment: Most stat softwares for regression produce a parameter table

Para.	Esti.	S.E.	t-value	p-value
β_0	\widehat{eta}_0	$S_{\widehat{eta}_0}$	$\widehat{eta}_0/S_{\widehat{eta}_0}$	2P(t(n-p) > t - value)
:	:	÷	:	÷
β_{p-1}	$\hat{\beta}_{p-1}$	$S_{\widehat{\beta}_{p-1}}$	$\widehat{\beta}_{p-1}/S_{\widehat{\beta}_{p-1}}$	2P(t(n-p) > t-value)

3. Scheffe's simultaneous CIs

(1) Extended Cauchy-Schwartz inequality

$$0 \le (x'y)^2 \le (x'Ax)(y'A^{-1}y)$$
 where $A > 0$.

Pf: With A > 0, in Cauchy-Schwartz inequality $0 \le (x'y)^2 \le (x'x)(y'y)$, replacing x by $A^{1/2}x$, and y by $A^{-1/2}y$ leads to the extended Cauchy-Schwartz inequality.

- (2) A lemma: With A > 0, $0 \le x'Ax \le c \Longrightarrow y'x \in \pm \sqrt{c(y'A^{-1}y)}$ for all $0 \ne y \in R^p$.
 - **Pf:** With A > 0 by the extended Cauchy-Schwartz inequality $0 \le x'Ax \le c \implies 0 \le \frac{(y'x)^2}{y'A^{-1}y} \le x'Ax \le c$ for all $0 \ne y \in R^p$ $\implies -\sqrt{c} \le \frac{y'x}{\sqrt{y'A^{-1}y}} \le \sqrt{c}$ for all $0 \ne y \in R^p$ $\implies y'x \in \pm \sqrt{c(y'A^{-1}y)}$ for all $0 \ne y \in R^p$.
- (3) Scheffe's simultaneous CIs

For $Y \sim N(X\beta, \sigma^2 \Sigma)$ where X has full column rank p, let $0 \neq l_i \in \mathbb{R}^p$, i = 1, 2, ... Then

$$l_i'\widehat{\beta} \pm \sqrt{p F_{\alpha}(p, n-p)} S_{l_i'\widehat{\beta}}, \ i = 1, 2, \dots$$

are simultaneous CIs for $l'_i\beta$, i = 1, 2, ..., with overall CC $1 - \alpha$.

$$\begin{split} \mathbf{Pf:} \ \text{With } \widehat{\beta} &= (\Sigma^{-1/2}X)^+ \Sigma^{-1/2}Y = (X'\Sigma^{-1}X)^{-1}X'\Sigma^{-1}Y \sim N\left(\beta, \, \sigma^2(X'\Sigma^{-1}X)^{-1}\right), \\ l'_i\widehat{\beta} \sim N\left(l'_i\beta, \, \sigma^2l'_i(X'\Sigma^{-1}X)^{-1}l_i\right). \\ \text{Thus } S^2_{l'_i\widehat{\beta}} &= \text{MSE } l'(X'\Sigma^{-1}X)^{-1}l_i = l'_iA^{-1}l_i \text{ where } A = \frac{X'\Sigma^{-1}X}{MSE} > 0. \\ \text{Let } D &= \left(\frac{(\beta-\widehat{\beta})'(X'\Sigma^{-1}X)(\beta-\widehat{\beta})}{MSE} \leq p \, F_\alpha(p, \, n-p)\right) = \left((\beta-\widehat{\beta})'A(\beta-\widehat{\beta}) \leq c\right) \text{ where } \\ c &= p \, F_\alpha(p, \, n-p). \text{ Then } P(D) = 1 - \alpha. \\ \text{Let } D_i &= \left(l'_i\beta \in l'_i\widehat{\beta} \pm \sqrt{p \, F_\alpha(p, \, n-p)} \, S_{l'_i\widehat{\beta}}\right) = \left(l'_i(\beta-\widehat{\beta}) \in \pm \sqrt{c, \, l'_iA^{-1}l_i}\right). \\ \text{By the extended C-S inequality,} \end{split}$$

$$(\beta - \widehat{\beta})' A(\beta - \widehat{\beta}) \le c \Longrightarrow l'_i(\beta - \widehat{\beta}) \in \sqrt{c, l'_i A^{-1} l_i} \text{ for all } i.$$

Thus $D \subset D_i$ for all *i*. So $D \subset D_1 \cap D_2 \cdots$. Hence $1 - \alpha = P(D) \leq P(D_1 \cap D_2 \cdots)$. **Ex2:** In $Y = X\beta + \epsilon$, $\epsilon \sim N(0, \sigma^2 \Sigma)$, $Y \in \mathbb{R}^8$, $\beta \in \mathbb{R}^4$ and $S_{\widehat{\beta}_1} = 0.9022$.

- (i) Find the width of a 95% CI for β_1 . With $\alpha = 0.05$, $\hat{\beta}_1 \pm t_{\alpha/2}(n-p)S_{\hat{\beta}_1}$ has width $W = 2t_{\alpha/2}(n-p)S_{\hat{\beta}_1} = 2t_{0.025}(4)S_{\hat{\beta}_1} = 2 \times 2.776 \times 0.9022 = 5.009.$
- (ii) Among the CIs for β_1 , β_2 and β_3 with overall CC 0.95 constructed by Bonferroni method, find the width of the CI for β_1 . With $\alpha = 0.05$ and k = 3, $\hat{\beta}_1 \pm t_{\alpha/(2k)}(n-p)S_{\hat{\beta}_1}$ has width $W = 2 t_{\alpha/(2k)}(n-p)S_{\hat{\beta}_1} = 2 t_{0.0083}(4) S_{\hat{\beta}_1} = 2 \times 3.961 \times 0.9022 = 7.1472.$
- (iii) Among the CIs for all $l'_i\beta$ with overall CC 0.95 constructed by Scheffe's method, find the width of the CI for β_1 . $\hat{\beta}_1 \pm \sqrt{pF_\alpha(p, n-p)} S_{\hat{\beta}_1}$ has width $W = 2\sqrt{pF_\alpha(p, n-p)} S_{\hat{\beta}_1} = 2\sqrt{4F_{0.05}(4, 4)} S_{\hat{\beta}_1} = 2 \times \sqrt{4 \times 6.39} \times 0.9022 = 9.1225$