

L21: Orthogonal factor models

1. Orthogonal factor models

(1) Definition and iff conditions

Factor model $X - \mu = LF + \epsilon$ is an orthogonal model if the columns of the loading matrix L is orthogonal, i.e., $L'L \in R^{q \times q}$ is a diagonal matrix. Thus

$$\begin{aligned} X - \mu = LF + \epsilon \text{ is an orthogonal model} &\iff L'L \text{ is a diagonal matrix} \\ &\iff L'L = \text{diag}(f_1^2, \dots, f_q^2). \end{aligned}$$

(2) An orthogonal model

If L in factor model $X - \mu = LF + \epsilon$ is estimated by $\widehat{L} = P_I \Lambda_I^{1/2}$ where $P_I \in R^{p \times q}$ and $\Lambda_I \in R^{q \times q}$ are from the EVD $S = P \Lambda P' = (P_I, P_{II}) \begin{pmatrix} \Lambda_I & 0 \\ 0 & \Lambda_{II} \end{pmatrix} (P_I, P_{II})'$ then $\widehat{L}'\widehat{L} = \Lambda_I^{1/2} P_I P_I' \Lambda_I = \Lambda_I = \text{diag}(\lambda_1, \dots, \lambda_q)$. So $X - \mu = LF + \epsilon$ is treated as an orthogonal model and $\widehat{f}_i^2 = \lambda_i, i = 1, \dots, q$.

(3) Distribution of estimated total variances in X

The estimated total variances in X is

$$s_1^2 + \dots + s_p^2 = \text{tr}(S) = \lambda_1 + \dots + \lambda_p.$$

The total contribution from F is

$$\sum_{i=1}^p \widehat{h}_i^2 = \sum_{j=1}^q \widehat{f}_j^2 = \text{tr}(\widehat{L}\widehat{L}') = \text{tr}(\widehat{L}'\widehat{L}) = \lambda_1 + \dots + \lambda_q.$$

The total contribution from ϵ is

$$\sum_{i=1}^p \widehat{\psi}_i = \sum_{i=1}^p (s_i^2 - \widehat{h}_i^2) = (\lambda_1 + \dots + \lambda_p) - (\lambda_1 + \dots + \lambda_q) = \lambda_{q+1} + \dots + \lambda_p.$$

Ex1: Suppose $X - \mu = LF + \epsilon$ is an orthogonal model. Convert this model to that for standardized $X, Z = V^{-1/2}(X - \mu) = (V^{-1/2}L)F + (V^{-1/2}\epsilon) = L_z F + \epsilon_z$.

Here $L_z = V^{-1/2}L$. So $L_z' L_z = L' V^{-1/2} V^{-1/2} L = L' V^{-1} L$ may not be a diagonal matrix.

Thus $Z = L_z F + \epsilon_z$ may not be an orthogonal model.

2. Orthogonal factor model for Z

(1) Orthogonal factor model for Z

If $Z = V^{-1/2}(X - \mu) = L_z F + \epsilon_z$ is an orthogonal factor model, then $\text{Cov}(Z) = \rho = L_z L_z' + \Psi_z$. Here ρ is the correlation matrix of $X, \Psi_z = \text{diag}(\psi_{z1}, \dots, \psi_{zp})$ and $L_z' L_z = \text{diag}(f_{1z}^2, \dots, f_{qz}^2)$.

(2) Estimation

Initially estimate ρ by R , the sample correlation matrix for X . $\rho_{ii} = 1 = r_{ii}, i = 1, \dots, p$.

By EVD $R = P \Lambda P' = (P_I, P_{II}) \begin{pmatrix} \Lambda_I & 0 \\ 0 & \Lambda_{II} \end{pmatrix} (P_I, P_{II})'$

Estimate the main part of $\rho = L_z L_z' + \Psi_z, L_z L_z'$, by the main part of $R, P_I \Lambda_I P_I'$, i.e., L_z is estimated by $P_I \Lambda_I^{1/2}$. Consequently $\widehat{h}_{zi}^2 = \widehat{l}_{i1}^2 + \dots + \widehat{l}_{iq}^2, i = 1, \dots, p$, are obtained.

Estimated $\psi_z = 1 - \widehat{h}_{zi}^2$ by $\widehat{\psi}_{zi} = 1 - \widehat{h}_{zi}^2, i = 1, \dots, p$.

$\widehat{L}_z' \widehat{L}_z = \Lambda_I^{1/2} P_I' P_I \Lambda_I^{1/2} = \Lambda_I$. So $\widehat{f}_{jz}^2 = \lambda_j, j = 1, \dots, q$.

(3) Distribution of estimated total variances in Z

The estimated total variances in Z is

$$r_{11} + \dots + r_{pp} = p = \text{tr}(R) = \lambda_1 + \dots + \lambda_p = p.$$

The total contribution from F is

$$\sum_{i=1}^p \widehat{h}_{zi}^2 = \sum_{j=1}^q \widehat{f}_{jz}^2 = \text{tr}(\widehat{L}_z \widehat{L}_z') = \text{tr}(\widehat{L}_z' \widehat{L}_z) = \lambda_1 + \dots + \lambda_q.$$

The total contribution from ϵ is

$$\sum_{i=1}^p \widehat{\psi}_{zi} = \sum_{i=1}^p (1 - \widehat{h}_{zi}^2) = (\lambda_1 + \dots + \lambda_p) - (\lambda_1 + \dots + \lambda_q) = \lambda_{q+1} + \dots + \lambda_p.$$

Ex2: Suppose $Z = V^{-1/2}(X - \mu) = L_z F + \epsilon_z$ is an orthogonal model.

Convert this model to that for $X, (X - \mu) = V^{1/2}Z = (V^{1/2}L_z)F + (V^{1/2}\epsilon_z) = LF + \epsilon$.

Here $L = V^{1/2}L_z$. So $L'L = L_z' V^{1/2} V^{1/2} L_z = L_z' V^{-1} L_z$ may not be a diagonal matrix.

Thus $X - \mu = LF + \epsilon$ may not be an orthogonal model.

3. Computations

(1) Analysis on orthogonal factor model for Z

```
data a;
  infile "D:\ex.txt";
  input x1 x2 x3 x4;
proc factor nfactor=2;
  var x1 x2 x3 x4;
run;
```

Four parts of output

- (i) Eigenvalues of correlation matrix of X . $\lambda_i, i = 1, \dots, 4$. $\lambda_1 + \dots + \lambda_4 = 4$.
 - (ii) Factor pattern ($\hat{L}_z = (P_1\sqrt{\lambda_1}, P_2\sqrt{\lambda_2}) \in R^{4 \times 2}$).
 - (iii) Variation explained by each factors ($\hat{f}_{1z}^2 = \lambda_1$ and $\hat{f}_{2z}^2 = \lambda_2$).
 - (iv) Communality ($\hat{h}_{zi}^2, i = 1, \dots, 4$. $\sum_i \hat{h}_{zi}^2 = \sum_j \hat{f}_{jz}^2 = \lambda_1 + \lambda_2$).
- (2) Directly entering S

In 9.8 on p531 $S = \begin{pmatrix} 1 & 0.4 & 0.9 \\ 0.4 & 1 & 0.7 \\ 0.9 & 0.7 & 1 \end{pmatrix}$

(i) Entering S

data a (type='cov'); _TYPE_='COV'; input _NAME_ \$ x1 x2 x3; datalines;	x1 1 0.4 0.9 x2 0.4 1 0.7 x3 0.9 0.7 1 ;
data a (type='cov'); _TYPE_='COV'; input _NAME_ \$ x1 x2 x3; datalines;	x1 1 0.4 0.9 x2 . 1 0.7 x3 . . 1 ;
data a (type='cov'); _TYPE_='COV'; input _NAME_ \$ x1 x2 x3; datalines;	x1 1 . . x2 0.4 1 . x3 0.9 0.7 1 ;

(ii) Analysis on orthogonal model for X

```
proc factor nfactor=1 cov;
  var x1 x2 x3;
run;
```

(iii) Analysis on orthogonal factor model for Z

```
proc factor nfactor=1;
  var x1 x2 x3;
run;
```