

## L14 Relation of test and confidence region

### 1. Confidence region and tests

Consider a population  $N(\mu, \Sigma)$ .

#### (1) Confidence region

$1 - \alpha$  confidence region for  $\mu \in R^p$  is the collection of all  $\mu \in R^p$  satisfying

$$(\mu - \bar{X})' \left( \frac{S}{n} \right)^{-1} (\mu - \bar{X}) \leq T_{\alpha}^2(p, n - 1)$$

#### (2) $\alpha$ -level Test

$H_0 : \mu = \mu_0$  vs  $H_a : \mu \neq \mu_0$   
 Test Statistic:  $T^2 = (\bar{X} - \mu_0)' \left( \frac{S}{n} \right)^{-1} (\bar{X} - \mu_0)$   
 Reject  $H_0$  if  $T^2 > T_{\alpha}^2(p, n - 1)$

#### (3) Test by $p$ -value

$H_0 : \mu = \mu_0$  versus  $H_a : \mu \neq \mu_0$   
 Test Statistic  $T^2 = (\bar{X} - \mu_0)' \left( \frac{S}{n} \right)^{-1} (\bar{X} - \mu_0)$   
 $p$ -value:  $P(T^2(p, n - 1) > T_{ob}^2)$

### 2. Relations

#### (1) Determining if $\mu_0$ in the $1 - \alpha$ confidence region for $\mu$

$$\begin{aligned} & \mu_0 \text{ is in a } 1 - \alpha \text{ confidence region for } \mu \\ \iff & \mu = \mu_0 \text{ satisfies } (\mu - \bar{X})' \left( \frac{S}{n} \right)^{-1} (\mu - \bar{X}) \leq T_{\alpha}^2(p, n - 1) \\ \iff & T^2 = (\bar{X} - \mu_0)' \left( \frac{S}{n} \right)^{-1} (\bar{X} - \mu_0) \leq T_{\alpha}^2(p, n - 1) \\ \iff & \alpha\text{-level test fails to reject } H_0 : \mu = \mu_0 \\ \iff & p\text{-value} > \alpha \end{aligned}$$

**Ex1:** To test  $H_0 : \mu = \begin{pmatrix} 6 \\ 9 \end{pmatrix}$  one ran SAS and obtained

Statistics	Value	F-value	Num DF	Den DF	Pr>F
Wilks' Lambda	0.7200	0.19	2	1	0.8485

Is  $\begin{pmatrix} 9 \\ 6 \end{pmatrix}$  in an 80% confidence region for  $\mu$ ?

Note that  $\begin{pmatrix} 6 \\ 9 \end{pmatrix}$  is in an 80% confidence region for  $\mu \iff p\text{-value} > 0.20$

But  $p\text{-value} = 0.8485 > 0.20$ . So yes.  $\begin{pmatrix} 6 \\ 9 \end{pmatrix}$  is in the 80% confidence region for  $\mu$ .

#### (2) Find smallest confidence coefficient such that $\mu_0$ is in the confidence region

$$\begin{aligned} \mu_0 \text{ is in a } 1 - \alpha \text{ confidence region for } \mu & \iff p\text{-value} > \alpha \\ & \iff 1 - \alpha > 1 - (p\text{-value}). \end{aligned}$$

**Ex2:** In Ex1 find the smallest confidence coefficient such that  $\begin{pmatrix} 6 \\ 9 \end{pmatrix}$  is in the confidence region.

With  $p\text{-value} = 0.8485$ ,

$$\begin{aligned} \mu_0 \text{ is in a } 1 - \alpha \text{ confidence region for } \mu & \iff p\text{-value} > \alpha \\ \iff & 1 - \alpha > 1 - (p\text{-value}) = 0.1515 = 15.15\% \end{aligned}$$

The smallest confidence coefficient such that  $\binom{6}{9}$  is in the confidence region is 15.15%.

### 3. Univariate two-sample problems

#### (1) Populations, samples and basic statistics

Populations:  $N(\mu_1, \sigma^2)$  and  $N(\mu_2, \sigma^2)$ .

Samples:  $X_1, \dots, X_{n_1}$  from Populations and  $Y_1, \dots, Y_{n_2}$  from population 2

Basic statistics:  $n_1, \bar{X}, \text{CSS}_1 = \sum_i (X_i - \bar{X})^2$  and  $s_1^2 = \frac{\text{CSS}_1}{n_1 - 1}$

$n_2, \bar{Y}, \text{CSS}_2 = \sum_i (Y_i - \bar{Y})^2$  and  $s_2^2 = \frac{\text{CSS}_2}{n_2 - 1}$

$n = n_1 + n_2, \text{CSS} = \text{CSS}_1 + \text{CSS}_2, s_p^2 = \frac{\text{CSS}}{n - 2} = \frac{n_1 - 1}{n - 2} s_1^2 + \frac{n_2 - 1}{n - 2} s_2^2.$

#### (2) Point estimators and confidence intervals

$\mu_1$  has MLE  $\bar{X}$  that is an UE;

$\mu_2$  has MLE  $\bar{Y}$  that is an UE;

$\sigma^2$  has MLE  $\frac{\text{CSS}}{n}$  and UE  $s_p^2$ .

$\left( \frac{\text{CSS}}{\chi_{\alpha/2}^2(n-2)}, \frac{\text{CSS}}{\chi_{1-\alpha/2}^2(n-2)} \right)$  is a  $1 - \alpha$  CI for  $\sigma^2$

$\bar{X} - \bar{Y} \pm t_{\alpha/2}(n - 2) \sqrt{s_p^2 \left( \frac{1}{n_1} + \frac{1}{n_2} \right)}$  is a  $1 - \alpha$  CI for  $\mu_1 - \mu_2$ .

#### (3) Tests

$H_0 : \mu_1 - \mu_2 = \delta_0$ vs $H_a : \mu_1 - \mu_2 \neq \delta_0$ Test statistic: $f = \frac{\bar{X} - \bar{Y} - \delta_0}{\sqrt{s_p^2 \left( \frac{1}{n_1} + \frac{1}{n_2} \right)}}$ Reject $H_0$ if $t < -t_{\alpha/2}(n - 2)$ or $t > t_{\alpha/2}(n - 2)$
$H_0 : \mu_1 - \mu_2 = \delta_0$ vs $H_a : \mu_1 - \mu_2 \neq \delta_0$ Test statistic: $f = \frac{\bar{X} - \bar{Y} - \delta_0}{\sqrt{s_p^2 \left( \frac{1}{n_1} + \frac{1}{n_2} \right)}}$ p-value: $2P(t(n - 2) >  t_{ob} )$

#### (4) Two multivariate populations

Two populations  $N_p(\mu_x, \Sigma)$  and  $N_p(\mu_y, \Sigma)$  can be expressed as

$$N_{p \times 2}(\mu, \Sigma, I_2) \text{ where } \mu = (\mu_x, \mu_y) \in R^{p \times 2}$$

which is equivalent to  $N \left( \begin{pmatrix} \mu_x \\ \mu_y \end{pmatrix}, \begin{pmatrix} \Sigma & 0 \\ 0 & \Sigma \end{pmatrix} \right)$ .

Two samples  $X \in R^{p \times n_1}$  from  $N_p(\mu_x, \Sigma)$  and  $Y \in R^{p \times n_2}$  from  $N_p(\mu_y, \Sigma)$  can be individually expressed as

$$X \sim N_{p \times n_1}(\mu_x \mathbf{1}'_{n_1}, \Sigma, I_{n_1}) \text{ and } Y \sim N_{p \times n_2}(\mu_y \mathbf{1}'_{n_2}, \Sigma, I_{n_2}),$$

or expressed together by

$$(X, Y) \sim N_{p \times n}(\mu J', \Sigma, I_n)$$

where  $n = n_1 + n_2, \mu = (\mu_x, \mu_y) \in R^{p \times 2}$  and  $J = \begin{pmatrix} \mathbf{1}_{n_1} & 0 \\ 0 & \mathbf{1}_{n_2} \end{pmatrix}$ . The expression is equivalent to

$$\begin{pmatrix} X_1 \\ \vdots \\ X_{n_1} \\ Y_1 \\ \vdots \\ Y_{n_2} \end{pmatrix} \sim N \left( \begin{pmatrix} \mu_x \\ \vdots \\ \mu_x \\ \mu_y \\ \vdots \\ \mu_y \end{pmatrix}, \begin{pmatrix} \Sigma & \cdots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \cdots & \Sigma \end{pmatrix} \right).$$